

On Bivariate Modeling of the COVID-19 Data with a New Type I Half-Logistic Inverse Weibull Distribution

Ahmed Elhassanein

Department of Mathematics, College of Science, University of Bisha, Bisha, Saudi Arabia

Department of Mathematics, Faculty of Science, Damanhour University, Damanhour, Egypt

Email: el_hassanein@yahoo.com

Article History

Received: 6 May, 2022


Revised: 11 July, 2022

Accepted: 15 August, 2022

Published: 23 August, 2022

Copyright © 2022 ARPG
& Author

This work is licensed under
the Creative Commons
Attribution International

 CC BY: Creative
Commons Attribution
License 4.0

Abstract

This manuscript presents a new univariate six parameters type I half-logistic inverse Weibull distribution. Explicit expressions for the quantile function, the moments, the moment generating function and the maximum likelihood estimators are formulated. Simulation is employed to investigate the goodness of fit and to discuss the behaviour of the new model. Competitive models are compared via real data. The univariate one is used as a base line to construct a bivariate one named bivariate six parameters type I half-logistic inverse Weibull distribution. Mathematical properties of the new bivariate distribution are investigated. The goodness of fit and the model performance are discussed via simulation. COVID-19 mortality data for Italy and Canada are treated as a bivariate random variable to prove the applicability of the new bivariate distribution.

Keywords: Bivariate distribution; Six parameters type I half-logistic inverse Weibull distribution; Maximum likelihood estimators; Bias; Standard error.

1. Introduction

Parametric probability distributions have a lot of applications especially in data analysis, statistical learning, and image processing [1-5]. Recently there has been a grating interest in formulating new parametric probability distributions. The inverse Weibull (IW) distribution is frequently considered in the literature. The suitable use of the IW model to describe the degeneration phenomena of mechanical components such as the dynamic components (pistons, crankshaft, etc.) of diesel engines has been discussed by Keller and Kamath [6]. It has been employed by Erto and Rapone [7] to model the times to break-down of the insulating fluid, subject to the action of constant tension. Akgul, *et al.* [8] considered the two-parameter IW model to investigate the wind speed. The three-parameter IW distribution has been constructed by De Gusmo, *et al.* [9]. Its properties have been discussed by Oluyede and Yang [10] and Jana and Bera [11]. Using the Marshall-Olkin method, another three-parameter IW model has been presented by Okasha, *et al.* [12]. The type I half-logistic two-parameter G family has been investigated by Cordeiro, *et al.* [13]. Alkarni, *et al.* [14] presented a new three parameter one as a generalized version. The univariate models can be applied only in the cases where there is an only one random variable or where there is a set of independent random variables.

They fail in the case where there is a set of dependent random variables. The problem of constructing bivariate models has a grate attention [15-28]. Here we aim: to introduce a new univariate six-parameter type I half logistic inverse Weibull (SPTIHLIW) distribution; to derive explicit mathematical expressions for its statistical quantities; to show its flexibility; to discuss the goodness of fit; to prove its superiority in comparison with a set of well-known models; to investigate its applicability to real data; to extend the univariate one to a bivariate one named bivariate six parameters type I half-logistic inverse Weibull (BISPTIHLIW) distribution; to derive the properties of the BISPTIHLIW distribution including, joint density function, joint cumulative function, conditional distributions, joint moments, hazard bivariate function and copula function; to investigate its performance; to emphasize the goodness of fit; to show the applicability of the BISPTIHLIW distribution for different types of data. As far as I know, the most of available distribution can only be applied for well-behaved data. The new model can be applied for ill-conditioned data including heavy tailed data. It has a joint probability density function with only one form with no singular parts. The pdf offers different shapes for different values of parameters. The hazard function has different shapes. It shows also good performance in terms of simulation study and application of real data. In addition, same algorithm was used to generate bivariate models with common properties including different shapes of the pdf, with singular part and no closed form of the maximum likelihood estimators. This paper uses a more general, easy, and different algorithm to present a bivariate distribution, with absolutely continuous pdf with new marginals. The

marginal are flexible with so favorable properties. Although the statistical quantities are completed they are in closed forms. The pdf shows different shapes and characteristics for different values of parameters and so its hazard function. The forthcoming part of the manuscript is organized as follows: The new univariate six-parameter type I half-logistic inverse Weibull (SPTIHLIW) distribution with its properties are constructed in Section 2. Section 3 is devoted to estimate the unknown parameters. Goodness of fit is discussed via simulation in section 4. Section 5 gives a real data application with comparison to competitive well-known models. The bivariate six-parameter type I half logistic inverse Weibull (BISTIHLIW) distribution is formulated in section 6. Estimation of unknown parameters and goodness of fit are presented in sections 7 and 8, respectively. Section 9 gives an application of the BISTIHLIW model. Concluding notes are given in section 10.

2. The SPTIHLIW Model

Definition 2.1. A one dimensional random variable Φ is said to follow the SPTIHLIW distribution if its cdf has the following form.

$$F_{\Phi}(\phi) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\vartheta} \tag{1}$$

where $\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta > 0$ and $\phi > 0$. It will be denoted by $\Phi \sim SPTIHLIW(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta)$

Proposition 2.2. For the cdf (1) the pdf is derived as

$$\begin{aligned} f_{\Phi}(\phi) &= 2\epsilon\varepsilon\zeta\eta\theta\vartheta\phi^{-\theta-1}e^{-\epsilon\phi^{-\epsilon}}(1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta-1} \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \\ &\times \frac{1}{(1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta})^2} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta-1} \\ &\times \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\vartheta-1} \end{aligned} \tag{2}$$

Proposition 2.3. An expansion of the pdf (2) can be obtained as follows.

$$f_{\Phi}(\phi) = \epsilon\varepsilon \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} \Omega_{j,k,l,m} \tag{3}$$

where $\Omega_{j,k,l,m} = m \binom{\zeta(k+l)}{m} \phi^{-\varepsilon-1} e^{-\varepsilon m \phi^{-\varepsilon}}$, and $0 < \phi < 1$.

Proposition 2.4. The survival function for the cdf (1) is computed as

$$S(\phi) = 1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\vartheta} \tag{4}$$

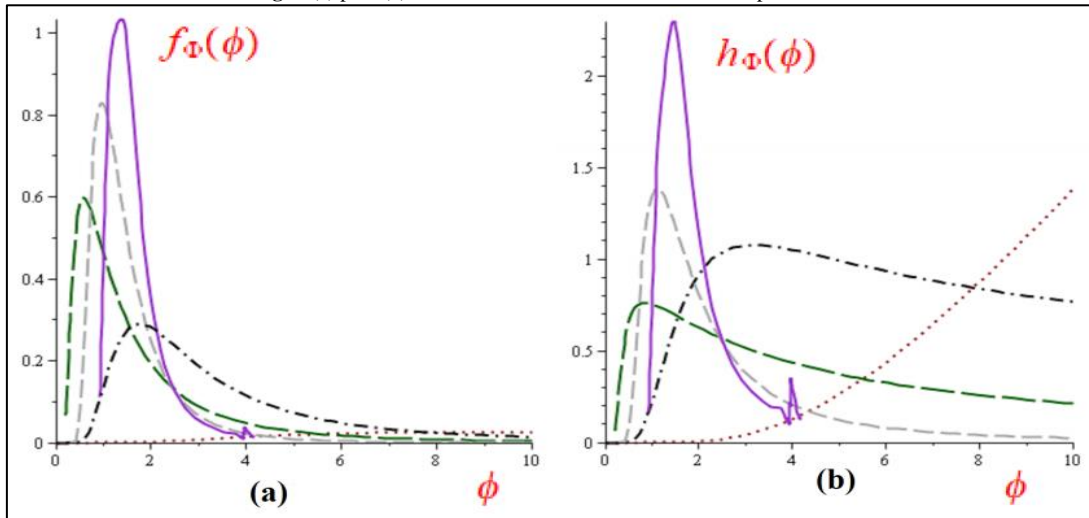
where $\phi > 0$:

Proposition 2.5. From (3) and (4) the hazard function is constructed as

$$\begin{aligned} h_{\Phi}(\phi) &= 2\epsilon\varepsilon\zeta\eta\theta\vartheta\phi^{-\theta-1}e^{-\epsilon\phi^{-\epsilon}}(1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta-1} \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \\ &\times \frac{1}{(1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta})^2} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta-1} \\ &\times \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\vartheta-1} \\ &\times \left[1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\vartheta} \right]^{-1} \end{aligned} \tag{5}$$

Figure 1, gives the pdf and the hazard function for different values of parameters. It shows the flexibility of the pdf and changes of the hazard function according to parameters.

Fig-1. (a) pdfs (b) hazard functions, for different values of parameters



Proposition 2.6. Expressions for the S^{th} moment and the moment generating function for the pdf (2) are respectively given by.

$$m_s = \epsilon^{\frac{s}{\epsilon}} \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}, \quad (6)$$

$$M(t) = \epsilon^{\frac{h}{\epsilon}} \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} \sum_{h=0}^{\infty} (-1)^{i+j+k-\eta j+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} \Psi_{j,k,l,m,h}, \quad (7)$$

where $F_{j,k,l,m} = m^{\frac{s}{\epsilon}} \binom{\zeta(k+l)}{m} \Gamma(1 - \frac{s}{\epsilon})$ and $\Psi_{j,k,l,m,h} = m^{\frac{h}{\epsilon}} \binom{\zeta(k+l)}{m} \Gamma(1 - \frac{h}{\epsilon}) \frac{t^h}{h!}$.

Proposition 2.7. The quantile function for the cdf (1) is

$$q(u) = \left\{ -\frac{1}{\epsilon} \ln \left(1 - \left[\frac{1 - \varrho(u)}{1 + \varrho(u)} \right]^{\frac{1}{\zeta}} \right) \right\}^{-\frac{1}{\epsilon}}, \quad (8)$$

where $\varrho(u) = (1 - (1 - u^{\frac{1}{\vartheta}})^{\frac{1}{\theta}})^{\frac{1}{\eta}}$ and $u \in (0, 1)$.

3. The SPTIHLIW Model: Estimation

Proposition 3.1. (I) The maximum likelihood function for the pdf (2) is computed as follows

$$\begin{aligned} H_{\Phi}(\epsilon, \vartheta, \zeta, \eta, \theta, \vartheta) &= n \ln(2\epsilon\vartheta\zeta\eta\theta\vartheta) - (\theta + 1) \sum_{i=1}^n \ln \phi_i - \epsilon \sum_{i=1}^n \phi_i^{-\epsilon} \\ &+ (\zeta - 1) \sum_{i=1}^n \ln(1 - e^{-\epsilon\phi_i^{-\epsilon}}) + (\eta - 1) \sum_{i=1}^n \ln\left(\frac{1 - (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}\right) \\ &- 2 \sum_{i=1}^n \ln(1 + (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}) + (\theta - 1) \sum_{i=1}^n \ln\left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}\right]^{\eta}\right) \\ &+ (\vartheta - 1) \sum_{i=1}^n \ln\left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon\phi_i^{-\epsilon}})^{\zeta}}\right]^{\eta}\right)^{\theta}\right) \end{aligned}$$

(II) The score function

$\Sigma = (\Sigma_{\epsilon}, \Sigma_{\vartheta}, \Sigma_{\zeta}, \Sigma_{\eta}, \Sigma_{\theta}, \Sigma_{\vartheta})'$, where $\Sigma_{\epsilon} = \Sigma_{\epsilon_1} + \Sigma_{\epsilon_2} + \Sigma_{\epsilon_3}$, $\Sigma_{\vartheta} = \Sigma_{\vartheta_1} + \Sigma_{\vartheta_2} + \Sigma_{\vartheta_3}$, $\Sigma_{\zeta} = \Sigma_{\zeta_1} + \Sigma_{\zeta_2} + \Sigma_{\zeta_3}$, $\Sigma_{\eta} = \Sigma_{\eta_1} + \Sigma_{\eta_2}$, is given by

$$\Sigma_{\epsilon_1} = \frac{n}{2\epsilon} - \sum_{i=1}^n \phi_i^{-\epsilon} + (\zeta - 1) \sum_{i=1}^n \frac{\phi_i^{-\epsilon} e^{-\epsilon\phi_i^{-\epsilon}}}{(1 - e^{-\epsilon\phi_i^{-\epsilon}})}$$

$$\begin{aligned} & \times \sum_{i=1}^n \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{2\zeta}} - 2\zeta \sum_{i=1}^n \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1}}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}, \\ \Sigma_{\varepsilon_2} &= 2\zeta\eta(\theta - 1) \sum_{i=1}^n \left\{ \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1}}{(1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right]^{\eta-1} \right. \\ & \times \left. \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^{-1} \right\}, \\ \Sigma_{\varepsilon_3} &= 2\zeta\eta\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1}}{(1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right]^{\eta-1} \right. \\ & \times \left. \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \right\}, \\ \Sigma_{\varepsilon_1} &= \frac{n}{2\varepsilon} + \varepsilon \sum_{i=1}^n \phi_i^{-\varepsilon} \ln \phi_i - \varepsilon(\zeta - 1) \sum_{i=1}^n \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} \ln \phi_i}{(1-e^{-\varepsilon\phi_i^{-\varepsilon}})} + 2\zeta\varepsilon(\eta - 1) \\ & \times \sum_{i=1}^n \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta \ln \phi_i}{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{2\zeta}} + 2\zeta\varepsilon \sum_{i=1}^n \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1} \ln \phi_i}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \\ \Sigma_{\varepsilon_2} &= 2\zeta\varepsilon\eta(\theta - 1) \sum_{i=1}^n \left\{ \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1} \ln \phi_i}{(1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right]^{\eta-1} \right. \\ & \times \left. \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^{-1} \right\}, \\ \Sigma_{\varepsilon_3} &= 2\zeta\eta\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \frac{\phi_i^{-\varepsilon} e^{-\varepsilon\phi_i^{-\varepsilon}} (1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{\zeta-1} \ln \phi_i}{(1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right]^{\eta-1} \right. \\ & \times \left. \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \right\}, \\ \Sigma_{\zeta_1} &= \frac{n}{2\zeta} + \sum_{i=1}^n \ln(1 - e^{-\varepsilon\phi_i^{-\varepsilon}}) - 2(\eta - 1) \sum_{i=1}^n \frac{(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta \ln(1-e^{-\varepsilon\phi_i^{-\varepsilon}})}{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^{2\zeta}} \\ & - 2 \sum_{i=1}^n \frac{(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta \ln(1-e^{-\varepsilon\phi_i^{-\varepsilon}})}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}, \\ \Sigma_{\zeta_2} &= 2\eta(\theta - 1) \sum_{i=1}^n \left\{ \frac{(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta \ln(1-e^{-\varepsilon\phi_i^{-\varepsilon}})}{(1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right]^{\eta-1} \right. \\ & \times \left. \left(1 - \left[\frac{1-(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta}{1+(1-e^{-\varepsilon\phi_i^{-\varepsilon}})^\zeta} \right] \eta \right)^{-1} \right\}, \end{aligned}$$

$$\begin{aligned} \Sigma_{\zeta_3} &= 2\eta\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \frac{(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta \ln(1-e^{-\epsilon\phi_i^{-\epsilon}})}{(1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right]^{\eta-1} \right. \\ &\times \left. \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \right\}, \\ \Sigma_{\eta_1} &= \frac{n}{2\eta} + \sum_{i=1}^n \ln\left(\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}\right) - (\theta - 1) \\ &\times \sum_{i=1}^n \left\{ \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \ln\left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^{-1} \right\} \\ \Sigma_{\eta_2} &= -\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \ln\left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \right. \\ &\times \left. \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \right\}, \\ \Sigma_{\theta} &= \frac{n}{2\theta} - \sum_{i=1}^n \ln \phi_i + \sum_{i=1}^n \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta\right) \\ &- (\vartheta - 1) \sum_{i=1}^n \left\{ \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \right. \\ &\times \left. \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \ln \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right\}, \\ \text{and } \Sigma_{\vartheta} &= \frac{n}{2\vartheta} + \sum_{i=1}^n \ln \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_i^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right). \end{aligned}$$

4. The SPTIHLIW Model: Simulation

To investigate the goodness of fit the function (8) is used to generate samples with size ($n = 20, 50, 100$). Three different random variables are treated: $\Phi_1 \sim SPTIHLIW(1.3, 1.7, 0.5, 0.9, 2.1, 2.0)$; $\Phi_2 \sim SPTIHLIW(1.8, 0.7, 1.5, 1.3, 0.8, 1.3)$; $\Phi_3 \sim SPTIHLIW(0.9, 2.3, 2.1, 2.0, 1.4, 0.7)$. The maximum likelihood method is employed to estimate the unknown parameters. The average estimates with bias and standard errors for Φ_1, Φ_2 , and Φ_3 are presented in Table 1, Table 2 and Table 3, respectively. We can observe the good performance the ML method and the goodness of fit by following the bias and standard error for estimates.

Table-1. Estimates for $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (1.3, 1.7, 0.5, 0.9, 2.1, 2.0)$

Sample Size	Parameter	Average Estimate	Bias	SE
n = 20	ϵ	1.515	0.215	0.245
	ε	2.481	0.781	0.296
	ζ	1.112	0.612	0.121
	η	1.009	0.509	0.203
	θ	1.914	-0.186	0.294
	ϑ	1.589	-0.411	0.211
n = 50	ϵ	1.449	0.149	0.126
	ε	2.081	0.318	0.139
	ζ	1.086	0.586	0.101
	η	1.237	0.337	0.166
	θ	1.826	-0.274	0.192
	ϑ	1.734	-0.266	0.135
n = 100	ϵ	1.207	-0.093	0.128
	ε	1.860	0.160	0.081
	ζ	0.968	0.468	0.092
	η	1.291	0.391	0.270
	θ	1.916	-0.184	0.136
	ϑ	1.733	-0.267	0.118

Table-2. Estimates for $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (1.8, 0.7, 1.5, 1.3, 0.8, 1.3)$.

Sample Size	Parameter	Average Estimate	Bias	SE
n = 20	ϵ	1.614	-0.186	0.146
	ε	1.017	0.317	0.175
	ζ	1.125	-0.475	0.126
	η	1.802	0.202	0.135
	θ	1.505	0.705	0.237
	ϑ	2.043	0.743	0.178
n = 50	ϵ	1.615	-0.185	0.130
	ε	0.869	0.169	0.049
	ζ	1.239	-0.361	0.093
	η	1.907	0.607	0.163
	θ	1.750	0.950	0.196
	ϑ	1.657	0.357	0.127
n = 100	ϵ	1.605	-0.195	0.106
	ε	0.8975	0.197	0.054
	ζ	1.118	-0.482	0.072
	η	1.476	0.176	0.141
	θ	1.700	0.900	0.229
	ϑ	1.502	0.202	0.130

Table-3. Estimates for $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (0.9, 2.3, 2.1, 2.0, 1.4, 0.7)$.

Sample Size	Parameter	Average Estimate	Bias	SE
n = 20	ϵ	1.612	0.712	0.238
	ε	3.190	0.890	0.211
	ζ	2.366	0.266	0.424
	η	2.227	0.227	0.431
	θ	2.558	1.158	0.835
	ϑ	1.383	0.683	0.299
n = 50	ϵ	1.668	0.768	0.132
	ε	2.922	0.692	0.171
	ζ	2.266	0.166	0.241
	η	1.183	-0.817	0.131
	θ	1.981	0.581	0.253
	ϑ	0.924	0.224	0.243
n = 100	ϵ	1.465	0.565	0.117
	ε	3.005	0.705	0.153
	ζ	2.020	-0.080	0.230
	η	1.819	-0.181	0.098
	θ	1.396	-0.004	0.146
	ϑ	0.782	0.082	0.170

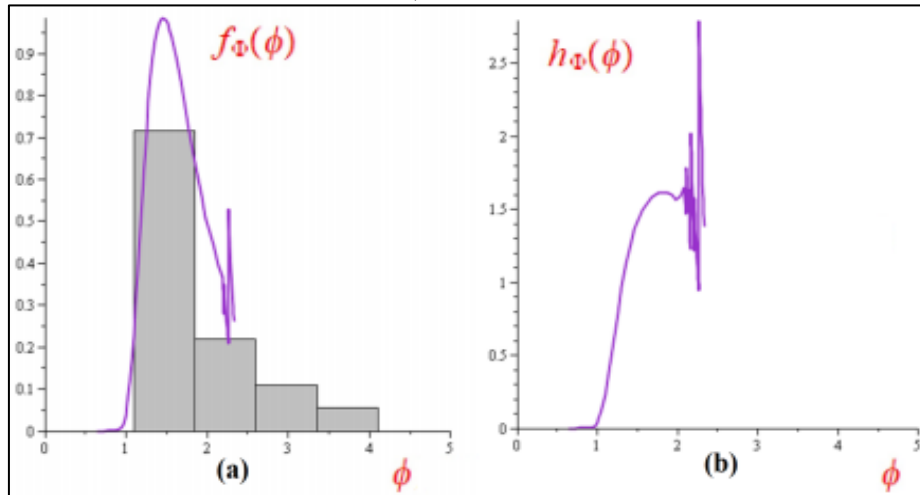
5. Relief Times Data

Analgesics are the drugs used to treat pain. The relief times data (in minutes) of 20 patients receiving an analgesic given by Gross and Clark [20] are modeled by the 10 new SPTIHLIW. The data are: 1.1, 1.4, 1.3, 1.7, 1.9, 1.8, 1.6, 2.2, 1.7, 2.7, 4.1, 1.8, 1.5, 1.2, 1.4, 3, 1.7, 2.3, 1.6, and 2. The ML estimates for the SPTIH-LIW model are $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (0:3892779, 8:9029837, 2:6088263, 0:1860415, 0:1630330, 8:8436032)$. The pdf and the hazard function are plotted in Fig. 2. The predefined competitive models are used for comparison, the BGIWGc model, BTW model, EHL-W, EG-W, and Kum-W, the results are displayed in Table 4, see [14, 29] for more details. From Table 4, the new SPTIHLIW shows good performance regarding to the value of AIC and BIC.

Table-4. The well-known models and the SPTIHLIW model with -L, AIC and BIC

Model	-L	AIC	BIC
SPTIHLIW	15.0227	42.0455	48.0199
BGIWGc	15.831	43.662	49.3639
BTW	16.5255	43.051	48.0297
EHL-W	17.113	42.226	46.2089
EG-W	17.486	42.972	46.9549
Kum-W	20.477	48.954	52.9369

Fig-2. The Ötted model for $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (0:3892779; 8:9029837; 2:6088263; 0:1860415; 0:1630330; 8:8436032)$



6. The BISPTIHLIW Model

The new bivariate six-parameter type I half logistic model is constructed using (1) as a base line distribution.

Definition 6.1. A two dimensional random variable (Φ_1, Φ_2) is said to follow the BISPTIHLIW distribution with parameters $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$, where $\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta > 0, -1 < \gamma_1 + \gamma_3 < 1, -1 < \gamma_2 + \gamma_3 < 1$ if its cumulative function is given by

$$F_{\Phi_1, \Phi_2}(\phi_1, \phi_2) = F_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2)F_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2) \tag{9}$$

where $F_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2) = 1 + (\gamma_1 + \gamma_3) \left(1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta} \right] \eta \right)^\theta \right\}^\vartheta \right) + (\gamma_2 + \gamma_3) \left(1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta} \right] \eta \right)^\theta \right\}^\vartheta \right)$, and

$F_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta} \right] \eta \right)^\theta \right\}^\vartheta \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta} \right] \eta \right)^\theta \right\}^\vartheta$, for $\phi_1 > 0$ and $\phi_2 > 0$. It will be denoted by $(\Phi_1, \Phi_2) \sim BISPTIHLIW(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$.

Proposition 6.2. The jpdf for $(\Phi_1, \Phi_2) \sim BISPTIHLIW(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ has the following form,

$$f_{\Phi_1, \Phi_2}(\phi_1, \phi_2) = f_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2)f_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2)f_{\Phi_1, \Phi_2}^3(\phi_1, \phi_2)f_{\Phi_1, \Phi_2}^4(\phi_1, \phi_2) \tag{10}$$

where $f_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2) = (2\epsilon\varepsilon\zeta\eta\theta\vartheta)^2(\phi_1\phi_2)^{-\theta-1}e^{-\epsilon(\phi_1^{-\varepsilon}+\phi_2^{-\varepsilon})}(1 - e^{-\epsilon\phi_1^{-\varepsilon}})^{\zeta-1}(1 - e^{-\epsilon\phi_2^{-\varepsilon}})^{\zeta-1} \frac{1 + \Upsilon(\gamma_1, \phi_1) + \Upsilon(\gamma_2, \phi_2)}{(1 + (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta)^2(1 + (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta)^2}$,

$$f_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2) = \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta} \right]^{\eta-1} \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta} \right]^{\eta-1},$$

$$f_{\Phi_1, \Phi_2}^3(\phi_1, \phi_2) = \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\varepsilon}})\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\varepsilon}})\zeta} \right] \eta \right)^{\theta-1},$$

$$f_{\Phi_1, \Phi_2}^4(\phi_1, \phi_2) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^{\vartheta-1} \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^{\vartheta-1},$$

and

$$\Upsilon(\gamma, x) = (\gamma + \gamma_3) \left(1 - 2 \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon x^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon x^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^\vartheta \right). \tag{11}$$

Proposition 6.3. The marginals for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ are derived as

$$F_{\Phi_1}(\phi_1) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^\vartheta \times \left\{ 1 + (\gamma_1 + \gamma_3) \left(1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^\vartheta \right) \right\}, \tag{12}$$

$$F_{\Phi_2}(\phi_2) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^\vartheta \times \left\{ 1 + (\gamma_2 + \gamma_3) \left(1 - \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^\vartheta \right) \right\}, \tag{13}$$

$$f_{\Phi_1}(\phi_1) = 2\epsilon\epsilon\zeta\eta\theta\vartheta\phi_1^{-\theta-1}e^{-\epsilon\phi_1^{-\epsilon}}(1 - e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta-1} \times \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^{\eta-1} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^\eta \right)^{\theta-1} \times \frac{1 + \Upsilon(\gamma_1, \phi_1)}{(1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta)^2} \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_1^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^{\vartheta-1} \tag{14}$$

and

$$f_{\Phi_2}(\phi_2) = 2\epsilon\epsilon\zeta\eta\theta\vartheta\phi_2^{-\theta-1}e^{-\epsilon\phi_2^{-\epsilon}}(1 - e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta-1} \times \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^{\eta-1} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^\eta \right)^{\theta-1} \times \frac{1 + \Upsilon(\gamma_2, \phi_2)}{(1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta)^2} \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_2^{-\epsilon}})\zeta} \right]^\eta \right)^\theta \right\}^{\vartheta-1} \tag{15}$$

Where $\Upsilon(\gamma, x)$ is given by (11).

Proposition 6.4. Expressions for the conditional densities for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ are

$$\begin{aligned}
 f_{\Phi_1/\Phi_2}(\phi_1/\phi_2) &= \frac{2\epsilon\epsilon\zeta\eta\theta\vartheta\phi_1^{-\theta-1}e^{-\epsilon\phi_1^{-\epsilon}}(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta})^2(1+\Upsilon(\gamma_2, \phi_2))} \left[\frac{1-(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \\
 &\quad \times (1+\Upsilon(\gamma_2, \phi_2) + \Upsilon(\gamma_1, \phi_1)) \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta-1} \\
 &\quad \times \left\{ 1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_1^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\theta-1} \tag{16}
 \end{aligned}$$

$$\begin{aligned}
 f_{\Phi_2/\Phi_1}(\phi_2/\phi_1) &= \frac{2\epsilon\epsilon\zeta\eta\theta\vartheta\phi_2^{-\theta-1}e^{-\epsilon\phi_2^{-\epsilon}}(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta})^2(1+\Upsilon(\gamma_1, \phi_1))} \left[\frac{1-(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \\
 &\quad \times (1+\Upsilon(\gamma_2, \phi_2) + \Upsilon(\gamma_1, \phi_1)) \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta-1} \\
 &\quad \times \left\{ 1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}}{1+(1-e^{-\epsilon\phi_2^{-\epsilon}})^{\zeta}} \right]^{\eta} \right)^{\theta} \right\}^{\theta-1} \tag{17}
 \end{aligned}$$

Where $\Upsilon(\gamma, x)$ is given by (11).

Proposition 6.5. The conditional moments for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ are calculated for (16) and (17) as

$$\begin{aligned}
 m_{\Phi_1/\Phi_2}^s(\phi_2) &= E(\Phi_1^s/\Phi_2 = \phi_2) = \left(1 + \frac{(\gamma_1 + \gamma_3)}{1 + \Upsilon(\gamma_2, \phi_2)} \right) \epsilon^{\frac{s}{\epsilon}} \\
 &\quad \times \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k-\eta j+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^s \\
 &\quad - \frac{2(\gamma_1 + \gamma_3)}{1 + \Upsilon(\gamma_2, \phi_2)} \epsilon^{\frac{s}{\epsilon}} \tag{18} \\
 &\quad \times \sum_{i_{1,2}=0}^{\vartheta} \sum_{j_{1,2}=0}^{\theta} \sum_{k_{1,2}=0}^{\eta i_{1,2}} \sum_{l_{1,2}=0}^{\infty} \sum_{m_{1,2}=1}^{\zeta(k_{1,2}+l_{1,2})} m_2 \xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}, m_{1,2}}^s
 \end{aligned}$$

and

$$\begin{aligned}
 m_{\Phi_2/\Phi_1}^s(\phi_1) &= E(\Phi_2^s/\Phi_1 = \phi_1) = \left(1 + \frac{(\gamma_2 + \gamma_3)}{1 + \Upsilon(\gamma_1, \phi_1)} \right) \epsilon^{\frac{s}{\epsilon}} \\
 &\quad \times \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k-\eta j+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^s \\
 &\quad + \frac{2(\gamma_2 + \gamma_3)}{1 + \Upsilon(\gamma_1, \phi_1)} \epsilon^{\frac{s}{\epsilon}} \tag{19} \\
 &\quad \times \sum_{i_{1,2}=0}^{\vartheta} \sum_{j_{1,2}=0}^{\theta} \sum_{k_{1,2}=0}^{\eta i_{1,2}} \sum_{l_{1,2}=0}^{\infty} \sum_{m_{1,2}=1}^{\zeta(k_{1,2}+l_{1,2})} m_2 \xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}, m_{1,2}}^s
 \end{aligned}$$

where

$$F_{j,k,l,m}^{\alpha} = m^{\alpha} \binom{\zeta(k+l)}{m} \Gamma\left(1 - \frac{\alpha}{\epsilon}\right), \tag{20}$$

$$\begin{aligned}
 \xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}}^{\alpha} &= (-1)^{i_1+i_2+j_1+j_2+k_1+k_2+l_1+l_2+m_1+m_2} \binom{\vartheta}{i_1} \binom{\vartheta}{i_2} \\
 &\quad \times \binom{\theta i_2}{j_2} \binom{\eta j_1}{k_1} \binom{\eta j_2}{k_2} \binom{\theta i_1}{j_1} \binom{\eta j_1+l-1}{l_1} \binom{\eta j_2+l-1}{l_2} \\
 &\quad \times \binom{\zeta(k_1+l)}{m_1} \binom{\zeta(k_2+l)}{m_2} \frac{\Gamma\left(1 - \frac{\alpha}{\epsilon}\right)}{(m_1 + m_2)^{1 - \frac{\alpha}{\epsilon}}}, \tag{21}
 \end{aligned}$$

Where $\Upsilon(\gamma, x)$ is given by (11).

Proposition 6.6. The joint moments for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ are

$$m_{r,s} = E(\Phi_1^r \Phi_2^s) = m_{r,s}^1 - m_{r,s}^2 - m_{r,s}^3 \tag{22}$$

where $m_{r,s}^1 =$

$$(1 + \gamma_1 + \gamma_2 + 2\gamma_3) \epsilon^{\frac{r}{\epsilon} + \frac{s}{\epsilon}} \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^r$$

$$\times \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^s,$$

$m_{r,s}^2 =$

$$2(\gamma_1 + \gamma_3) \epsilon^{\frac{r}{\epsilon} + \frac{s}{\epsilon}} \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^s$$

$$\times \sum_{i_{1,2}=0}^{\vartheta} \sum_{j_{1,2}=0}^{\theta i_{1,2}} \sum_{k_{1,2}=0}^{\eta j_{1,2}} \sum_{l_{1,2}=1}^{\infty} \sum_{m_{1,2}=1}^{\zeta(k_{1,2}+l_{1,2})} m_2 \xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}, m_{1,2}}^r,$$

$m_{r,s}^3 =$

$$2(\gamma_2 + \gamma_3) \epsilon^{\frac{r}{\epsilon} + \frac{s}{\epsilon}} \sum_{i=0}^{\vartheta} \sum_{j=0}^{\theta i} \sum_{k=0}^{\eta j} \sum_{l=0}^{\infty} \sum_{m=1}^{\zeta(k+l-\eta j)} (-1)^{i+j+k+l+m} \binom{\vartheta}{i} \binom{\theta i}{j} \binom{\eta j}{k} \binom{\eta j+l-1}{l} F_{j,k,l,m}^r$$

$$\times \sum_{i_{1,2}=0}^{\vartheta} \sum_{j_{1,2}=0}^{\theta i_{1,2}} \sum_{k_{1,2}=0}^{\eta j_{1,2}} \sum_{l_{1,2}=1}^{\infty} \sum_{m_{1,2}=1}^{\zeta(k_{1,2}+l_{1,2})} m_2 \xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}, m_{1,2}}^s,$$

$F_{j,k,l,m}^\alpha$, and $\xi_{i_{1,2}, j_{1,2}, k_{1,2}, l_{1,2}, m_{1,2}}^\alpha$ are given by (20) and (21), respectively.

Proposition 6.7. The bivariate reliability function for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$, is

$$r(\phi_1, \phi_2) = \frac{(\Upsilon(\gamma_1, \phi_1) + \gamma_1 + \gamma_3)(\Upsilon(\gamma_2, \phi_2) + \gamma_2 + \gamma_3)}{8(\gamma_1 + \gamma_3)}$$

$$\times \frac{(\Upsilon(\gamma_1, \phi_1) + \Upsilon(\gamma_2, \phi_2) - \gamma_1 - \gamma_2 - 2\gamma_3 + 2)}{(\gamma_2 + \gamma_3)} \tag{23}$$

Proposition 6.8. The bivariate hazard rate function for

$(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$, is

$$h_{\Phi_1, \Phi_2}(\phi_1, \phi_2) = h_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2) h_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2) h_{\Phi_1, \Phi_2}^3(\phi_1, \phi_2) h_{\Phi_1, \Phi_2}^4(\phi_1, \phi_2) \tag{24}$$

where,

$$h_{\Phi_1, \Phi_2}^1(\phi_1, \phi_2) = \frac{32(\gamma_1 + \gamma_3)(\gamma_2 + \gamma_3)(\epsilon \zeta \eta \theta \vartheta)^2 (\phi_1 \phi_2)^{-\theta-1} e^{-\epsilon(\phi_1^{-\epsilon} + \phi_2^{-\epsilon})} (1 - e^{-\epsilon \phi_1^{-\epsilon}})^{\zeta-1} (1 - e^{-\epsilon \phi_2^{-\epsilon}})^{\zeta-1}}{(1 + (1 - e^{-\epsilon \phi_1^{-\epsilon}})^{\zeta})^2 (1 + (1 - e^{-\epsilon \phi_2^{-\epsilon}})^{\zeta})^2} \frac{1}{(\Upsilon(\gamma_1, \phi_1) + \gamma_1 + \gamma_3)(\Upsilon(\gamma_2, \phi_2) + \gamma_2 + \gamma_3)},$$

$$h_{\Phi_1, \Phi_2}^2(\phi_1, \phi_2) = \left[\frac{1 - (1 - e^{-\epsilon \phi_1^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_1^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \left[\frac{1 - (1 - e^{-\epsilon \phi_2^{-\epsilon}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_2^{-\epsilon}})^{\zeta}} \right]^{\eta-1} \frac{1}{(\Upsilon(\gamma_1, \phi_1) + \Upsilon(\gamma_2, \phi_2) - \gamma_1 - \gamma_2 - 2\gamma_3 + 2)},$$

$$h_{\Phi_1, \Phi_2}^3(\phi_1, \phi_2) = \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_1^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_1^{-\zeta}})^{\zeta}} \right] \eta \right)^{\theta-1} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_2^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_2^{-\zeta}})^{\zeta}} \right] \eta \right)^{\theta-1},$$

and

$$h_{\Phi_1, \Phi_2}^4(\phi_1, \phi_2) = \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_1^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_1^{-\zeta}})^{\zeta}} \right] \eta \right)^{\theta} \right\}^{\vartheta-1} \left\{ 1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_2^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_2^{-\zeta}})^{\zeta}} \right] \eta \right)^{\theta} \right\}^{\vartheta-1}$$

Proposition 6.9. The copula function for $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$, is

$$c(x, y) = \frac{1 + \Upsilon(\gamma_1, x) + \Upsilon(\gamma_2, y)}{(1 + \Upsilon(\gamma_1, x))(1 + \Upsilon(\gamma_2, y))}, \tag{25}$$

where $\Upsilon(\epsilon, x)$ is given by (11), [30, 31]

7. The BISPTIHLIW Model: Estimation

Proposition 7.1. Let $(\Phi_{11}, \Phi_{21}), (\Phi_{12}, \Phi_{22}), \dots, (\Phi_{1n}, \Phi_{2n})$ be a random sample of size n from a random variable $(\Phi_1, \Phi_2) \sim \text{BISPTIHLIW}(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$.

Then:

(I) The maximum log-likelihood function is given by $H_{\Phi_1, \Phi_2}(\theta) = H_1(\theta) + H_2(\theta) + H_3(\theta) + H_4(\theta)$, where $\theta = (\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$,

$$H_1(\theta) = 2n \ln(2\epsilon\epsilon\zeta\eta\theta\vartheta) - (\theta + 1) \sum_{i=1}^n (\ln \phi_{1i} + \ln \phi_{2i}) - \epsilon \sum_{i=1}^n (\phi_{1i}^{-\zeta} + \phi_{2i}^{-\zeta})$$

$$+ (\zeta - 1) \sum_{i=1}^n \left(\ln(1 - e^{-\epsilon \phi_{1i}^{-\zeta}}) + \ln(1 - e^{-\epsilon \phi_{2i}^{-\zeta}}) \right),$$

$$H_2(\theta) = (\eta - 1) \sum_{i=1}^n \left(\ln \left(\frac{1 - (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}} \right) + \ln \left(\frac{1 - (1 - e^{-\epsilon \phi_{2i}^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_{2i}^{-\zeta}})^{\zeta}} \right) \right)$$

$$- 2 \sum_{i=1}^n \left(\ln(1 + (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}) + \ln(1 + (1 - e^{-\epsilon \phi_{2i}^{-\zeta}})^{\zeta}) \right),$$

$$H_3(\theta) = (\theta - 1) \sum_{i=1}^n \left\{ \ln \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}} \right] \eta \right) + \ln \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_{2i}^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_{2i}^{-\zeta}})^{\zeta}} \right] \eta \right) \right\}$$

$$+ \sum_{i=1}^n \ln(1 + \Upsilon(\gamma_1, \phi_{1i}) + \Upsilon(\gamma_2, \phi_{2i})),$$

$$\text{and } H_4(\theta) = (\vartheta - 1) \sum_{i=1}^n \left\{ \ln \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}}{1 + (1 - e^{-\epsilon \phi_{1i}^{-\zeta}})^{\zeta}} \right] \eta \right)^{\theta} \right) \right\}$$

$$+ \ln\left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_2^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_2^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)\right\}.$$

(II) The score function $F = (F_\epsilon, F_e, F_\zeta, F_\eta, F_\theta, F_\vartheta, F_{\gamma_1}, F_{\gamma_2}, F_{\gamma_3})'$ is given by $F_\epsilon = F_{\epsilon_1} + F_{\epsilon_2} + F_{\epsilon_3} + F_{\epsilon_4}$, $F_e = F_{e_1} + F_{e_2} + F_{e_3} + F_{e_4}$, $F_\zeta = F_{\zeta_1} + F_{\zeta_2} + F_{\zeta_3} + F_{\zeta_4}$, $F_\eta = F_{\eta_1} + F_{\eta_2} + F_{\eta_3}$, and $F_\theta = F_{\theta_1} + F_{\theta_2} + F_{\theta_3}$, $F_\vartheta = F_{\vartheta_1} + F_{\vartheta_2}$, F_{γ_1} , F_{γ_2} , and F_{γ_3} , where

$$F_{\epsilon_1} = \frac{n}{\epsilon} - \sum_{i=1}^n (\phi_{1i}^{-\epsilon} + \phi_{2i}^{-\epsilon}) + (\zeta - 1) \sum_{i=1}^n \left(\frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}}}{(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}}}{(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})} \right)$$

$$- 2\zeta(\eta - 1) \sum_{i=1}^n \left(\frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{2\zeta}} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{2\zeta}} \right)$$

$$- 2\zeta \sum_{i=1}^n \left(\frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1}}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1}}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right),$$

$$F_{\epsilon_2} = 2\zeta\eta(\theta - 1) \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta - 1 \right.$$

$$\times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta \right)^{-1} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta)^2}$$

$$\left. \times \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right]^\eta - 1 \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right]^\eta \right)^{-1} \right\},$$

$$F_{\epsilon_3} = 2\zeta\eta\theta\vartheta \sum_{i=1}^n \frac{1}{1+Y(\gamma_1, \phi_{1i})+Y(\gamma_2, \phi_{2i})} \left\{ \frac{(\gamma_1+\gamma_3)\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta)^2} \right.$$

$$\times \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta - 1 \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta \right)^\theta \right)^{\theta-1}$$

$$+ \frac{(\gamma_2+\gamma_3)\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right]^\eta - 1 \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right]^\eta \right)^\theta \right)^{\theta-1}$$

$$\left. \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right]^\eta \right)^{\theta-1} \right\},$$

$$F_{\epsilon_4} = 2\zeta\eta\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right]^\eta - 1 \right.$$

$$\begin{aligned} & \times \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \\ & + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta^{-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \\ & \times \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1} \Bigg\}, \\ F_{\epsilon_1} &= \frac{n}{\epsilon} + \epsilon \sum_{i=1}^n (\phi_{1i}^{-\epsilon} \ln \phi_{1i} + \phi_{2i}^{-\epsilon} \ln \phi_{2i}) + \epsilon(\zeta-1) \sum_{i=1}^n \left(\frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} \ln \phi_{1i}}{(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} \ln \phi_{2i}}{(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})} \right), \\ F_{\epsilon_2} &= 2\epsilon\zeta(\eta-1) \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta} \ln \phi_{1i}}{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{2\zeta}} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta} \ln \phi_{2i}}{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{2\zeta}} \right\} \\ & + 2\epsilon\zeta \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1} \ln \phi_{1i}}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1} \ln \phi_{2i}}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right\}, \\ F_{\epsilon_3} &= 2\epsilon\zeta\eta(\theta-1) \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1} \ln \phi_{1i}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1} \right. \\ & \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{-1} + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1} \ln \phi_{2i}}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta^{-1} \\ & \left. \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{-1} \right\}, \\ F_{\epsilon_4} &= 2\zeta\eta\theta\vartheta \sum_{i=1}^n \left\{ \frac{(\gamma_1+\gamma_3)\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1} \ln \phi_{1i}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1} \right. \\ & \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1} \\ & + \frac{(\gamma_2+\gamma_3)\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^{\zeta-1} \ln \phi_{2i}}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta^{-1} \\ & \left. \times \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \right\} \frac{1}{1+\Upsilon(\gamma_1, \phi_{1i}) + F(\gamma_2, \phi_{2i})}, \\ F_{\epsilon_5} &= 2\zeta\eta\theta(\vartheta-1) \sum_{i=1}^n \left\{ \frac{\phi_{1i}^{-\epsilon} e^{-\epsilon\phi_{1i}^{-\epsilon}} (1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^{\zeta-1}}{(1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1} \right. \end{aligned}$$

$$\begin{aligned} & \times \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \\ & + \frac{\phi_{2i}^{-\epsilon} e^{-\epsilon\phi_{2i}^{-\epsilon}} (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta^{-1} \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \\ & \times \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1} \Bigg\}, \end{aligned}$$

$$F_{\zeta_1} = \frac{n}{\zeta} + \sum_{i=1}^n \left(\ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}}) + \ln(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}}) \right)$$

$$- 2(\eta - 1) \sum_{i=1}^n \left\{ \frac{(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})}{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})^{2\zeta}} + \frac{(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})}{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})^{2\zeta}} \right\},$$

$$F_{\zeta_2} = -2 \sum_{i=1}^n \left\{ \frac{(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} + \frac{(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right\},$$

$$\begin{aligned} F_{\zeta_3} &= 2\eta(\theta - 1) \sum_{i=1}^n \left\{ \frac{(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}}) \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{-1} \right. \\ & \left. + \frac{(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})^\gamma \ln(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}}) \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \theta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2} \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{-1} \right\}, \end{aligned}$$

$$F_{\zeta_4} = 2\eta\theta(\vartheta - 1) \sum_{i=1}^n \left\{ \frac{(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta \ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}}) \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \right.$$

$$\times \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1}$$

$$+ \frac{(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta \ln(1 - e^{-\epsilon\phi_{2i}^{-\epsilon}}) \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta)^2}$$

$$\left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{-1} \Bigg\},$$

$$F_{\zeta_5} = 2\eta\theta\vartheta \sum_{i=1}^n \left\{ (\gamma_1 + \gamma_3) \frac{(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta \ln(1 - e^{-\epsilon\phi_{1i}^{-\epsilon}}) \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta^{-1}}{(1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta)^2} \right.$$

$$\times \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1 - (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1 + (1 - e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1}$$

$$\begin{aligned}
 & +(\gamma_2 + \gamma_3) \frac{(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta \ln(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})}{(1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta)^2} \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta^{-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \\
 & \left. \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1} \right\} \frac{1}{1+\Upsilon(\gamma_1, \phi_{1i}) + \Upsilon(\gamma_2, \phi_{2i})}, \\
 F_{\eta_1} & = \frac{n}{\eta} + \sum_{i=1}^n \left(\ln \left(\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right) + \ln \left(\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right) \right), \\
 F_{\eta_2} & = -(\theta - 1) \sum_{i=1}^n \left\{ \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right)^{-1} \right. \\
 & \left. + \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^{-1} \right\}, \\
 F_{\eta_3} & = \theta \vartheta \sum_{i=1}^n \left\{ \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \right. \\
 & \times \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} + \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \\
 & \left. \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{-1} \right\}, \\
 F_{\eta_4} & = 2\theta \vartheta \sum_{i=1}^n \left\{ (\gamma_1 + \gamma_3) \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \right. \\
 & \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1} \\
 & \left. + (\gamma_2 + \gamma_3) \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \ln \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^{\theta-1} \right. \\
 & \times \left. \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})^\zeta} \right] \eta \right)^\theta \right)^{\vartheta-1} \right\} \frac{1}{1+\Upsilon(\gamma_1, \phi_{1i}) + \Upsilon(\gamma_2, \phi_{2i})}, \\
 F_{\theta_1} & = \frac{n}{\theta} - \sum_{i=1}^n (\ln \phi_{1i} + \ln \phi_{2i}) + \sum_{i=1}^n \left\{ \ln \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})^\zeta} \right] \eta \right) \right.
 \end{aligned}$$

$$+ \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)\Bigg\},$$

$$F_{\theta_2} = 2(\vartheta - 1) \sum_{i=1}^n \left\{ \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^{-1} \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta \right. \\ \times \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right) + \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^{-1} \\ \left. \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)\right\},$$

$$F_{\theta_3} = 2\vartheta \sum_{i=1}^n \left\{ (\gamma_1 + \gamma_3) \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^{\vartheta-1} \right. \\ \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right) \\ \left. + (\gamma_2 + \gamma_3) \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^{\vartheta-1} \right. \\ \left. \times \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta \ln\left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)\right\} \frac{1}{1+\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})},$$

$$F_{\theta_1} = \sum_{i=1}^n \left\{ \ln\left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right) + \ln\left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right) \right\},$$

$$F_{\theta_2} = 2 \sum_{i=1}^n \left\{ (\gamma_1 + \gamma_3) \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^\vartheta \right. \\ \times \ln\left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{1i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right) + (\gamma_2 + \gamma_3) \left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)^\vartheta \\ \left. \times \ln\left(1 - \left(1 - \left[\frac{1-(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}{1+(1-e^{-\epsilon\phi_{2i}^{-\epsilon}})\zeta}\right]\eta\right)^\theta\right)\right\} \frac{1}{1+\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})},$$

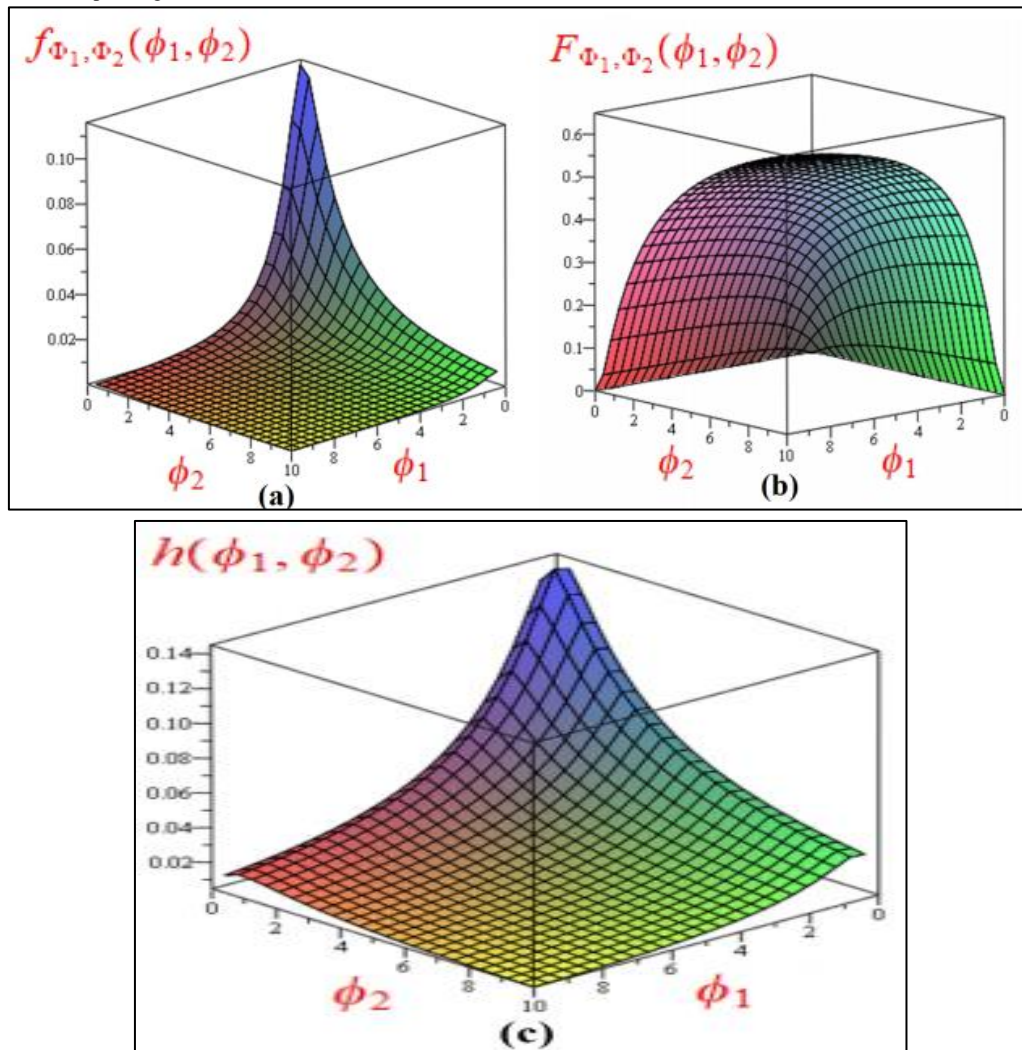
$$F_{\gamma_1} = 2 \sum_{i=1}^n \frac{\Upsilon(\gamma_1, \phi_{1i})}{1+\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})}, \quad F_{\gamma_2} = 2 \sum_{i=1}^n \frac{\Upsilon(\gamma_2, \phi_{2i})}{1+\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})},$$

And $F_{\gamma_3} = 2 \sum_{i=1}^n \frac{\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})}{1+\Upsilon(\gamma_1, \phi_{1i})+\Upsilon(\gamma_2, \phi_{2i})}$, where $\Upsilon(\gamma, x)$ is given by (11).

8. The BISPTIHLIW Model: Simulation

Here we discuss numerically the characteristics of the new BISPTIHLIW distribution then use simulation to investigate the goodness of fit. We consider three different random variables. The first one is the BISPTIHLIW (1.3, 0.9, 2.5, 0.9, 0.3, 0.8, 0.2, 0.9, 0.6) whose jpdf, jcdf, marginals, and copula are plotted in Fig. 3. We can observe the right skewness and unimodality of the jpdf and marginals in Fig. 3 (a), 3(f), and 3(g). The jcdf \rightarrow 0.6, Fig. 3 (b), and cdfs \rightarrow 0.5, Figs. 3 (g) and 3(i), where $(\Phi_1, \Phi_2) \rightarrow (10, 10)$, that suggests applying this model for the data with measurements scale in the interval $(0,20) \times (0,20)$. The hazard function is a decreasing function that approaches zero where $(\Phi_1, \Phi_2) \rightarrow (10,10)$, Fig. 3 (c). The second one is the BISPTIHLIW (2.1, 1.1, 0.7, 1.2, 0.9, 1.2, 0.6, 0.4, 0.8), its corresponding functions are displayed in Fig. 4. In this case the jcdf \rightarrow 0.2, Fig. 4 (b), and cdfs \rightarrow 0.9, Figs. 4 (g) and 4(i), where $(\Phi_1, \Phi_2) \rightarrow (10, 10)$, that suggests applying this model for the data with measurements scale in the interval $(0,80) \times (0,80)$: The hazard function changes its behaviour from increasing to decreasing for $(\phi_1, \phi_2) \in (10,10)$, Fig. 4 (c). The right skewness and unimodality will be clear for higher measurements scale, Figs. 4 (a), 4(f), 4(h). The third one is the BISPTIHLIW (0.9, 1.7, 1.9, 1.6, 1.3, 1.6, 0.6, 0.7, 0.8) Its related functions are given in Fig. 5. The jcdf \rightarrow 1:0, Fig. 3 (b), and cdfs \rightarrow 1.0, Figs. 3 (g) and 3(i), where $(\Phi_1, \Phi_2) \rightarrow (4, 4)$, that suggests applying this model for the data with measurements scale in the interval $(0,4) \times (0, 4)$: Also grantee the applicability of the new model of ratio data. The hazard function changes its behaviour from increasing to decreasing for $(\phi_1, \phi_2) \in (10, 10)$, Fig. 5 (c). Comparison between Figs. 3(d), 3(e), 4(d), 4(e) and 5(d), 5(e) shows the different levels of correlation between different set of variables. The goodness of fit is investigated via Monte Carlo simulation. Samples are generated with sizes $(n=30, n=50, n=120)$ for different three sets of parameters. The average estimates with bias and standard errors are displayed in Table 5, Table 6, and Table 7. Using bias and standard errors as a comparison criteria, we can see that the maximum likelihood method gives good estimates for the unknown parameters. We consider $\gamma_1, \gamma_2, \text{ and } \gamma_3$ as chosen parameters.

Fig-3. Functions corresponding to $(\epsilon, \epsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3) = (1.3, 0.9, 2.5, 0.9, 0.3, 0.8, 0.2, 0.9, 0.6)$.



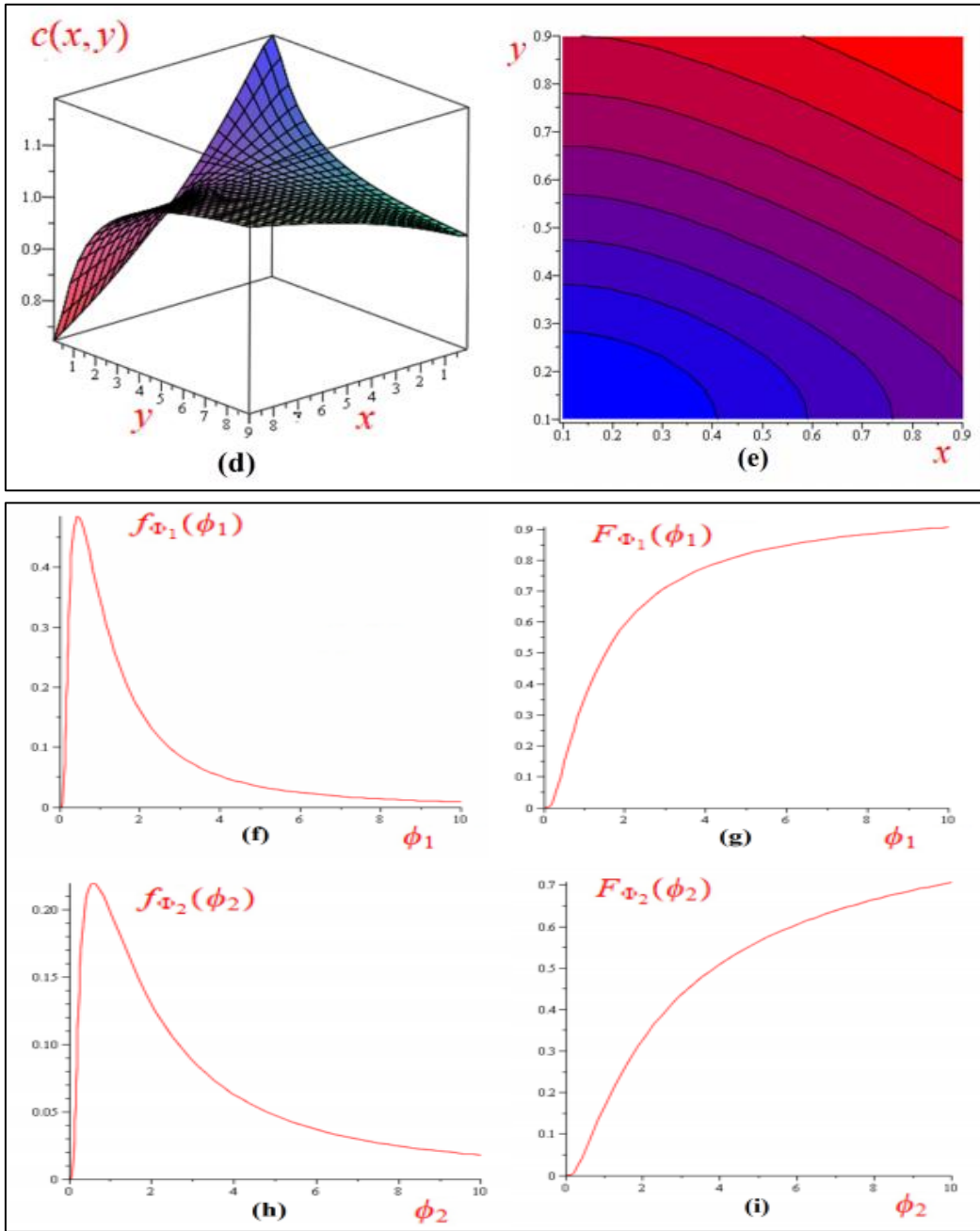
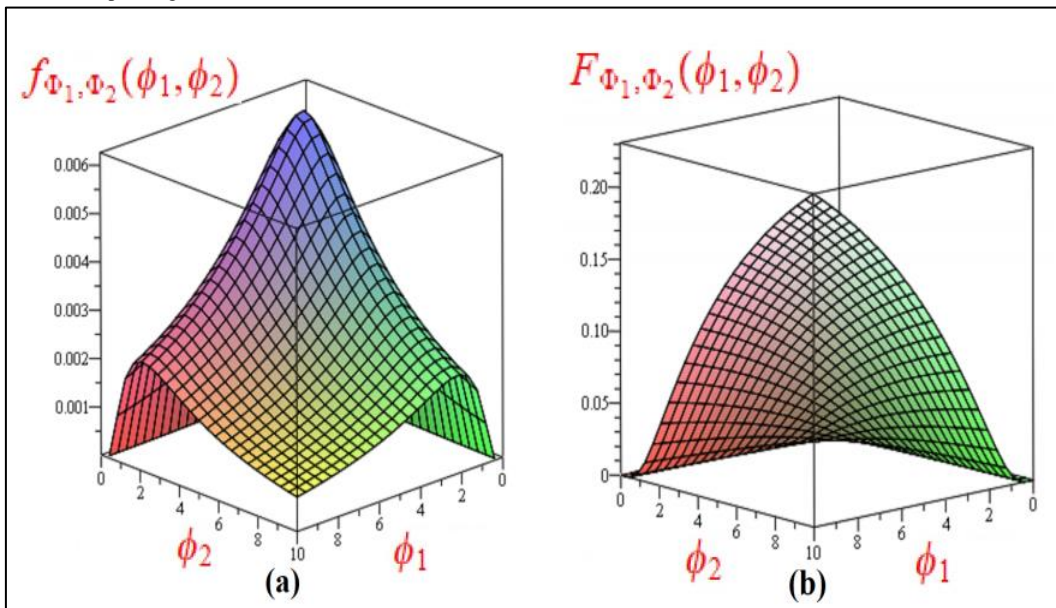


Fig-4. Functions corresponding to $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3) = (2:1; 1:1; 0:7; 1:2; 0:9; 1:2; 0:6; 0:4; 0:8)$



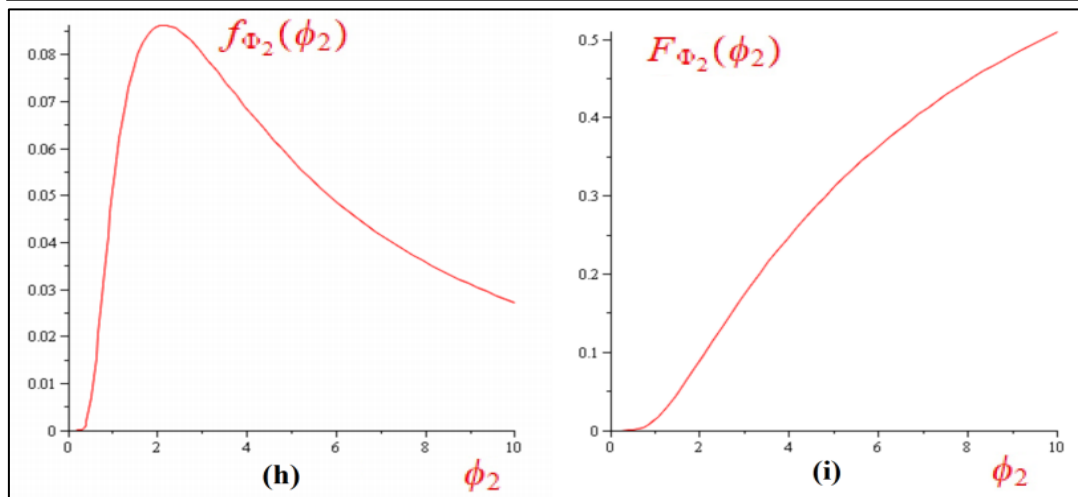
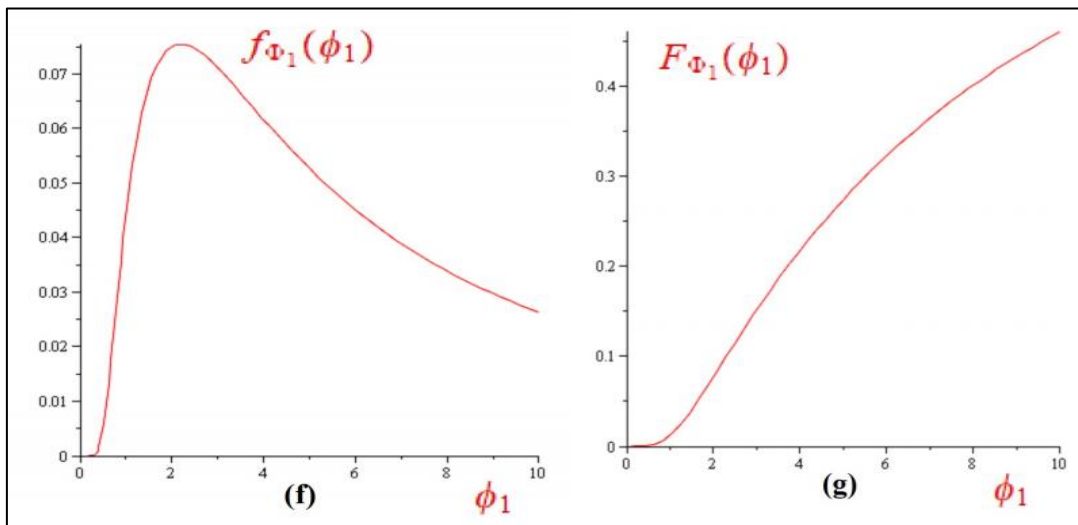
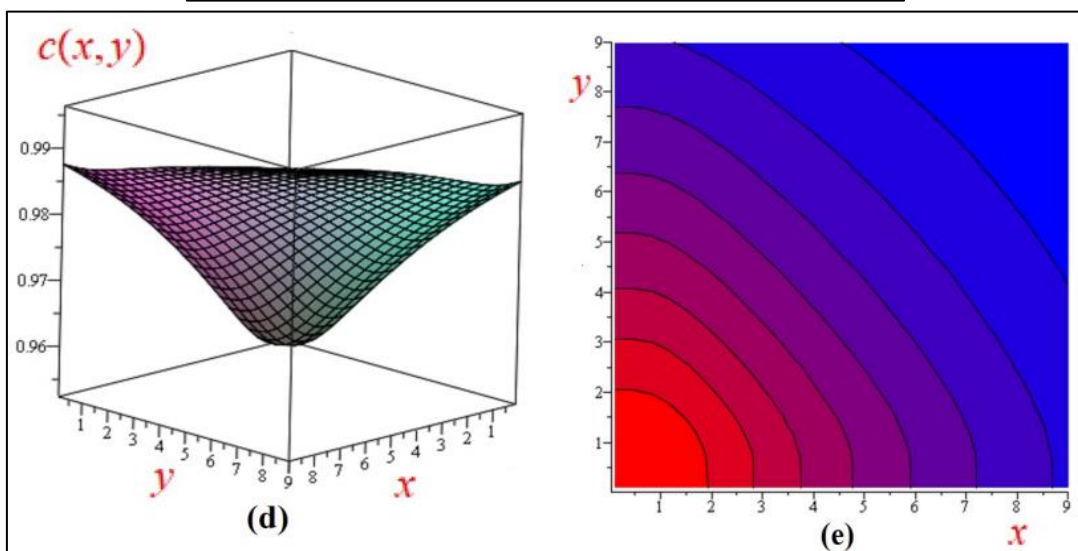
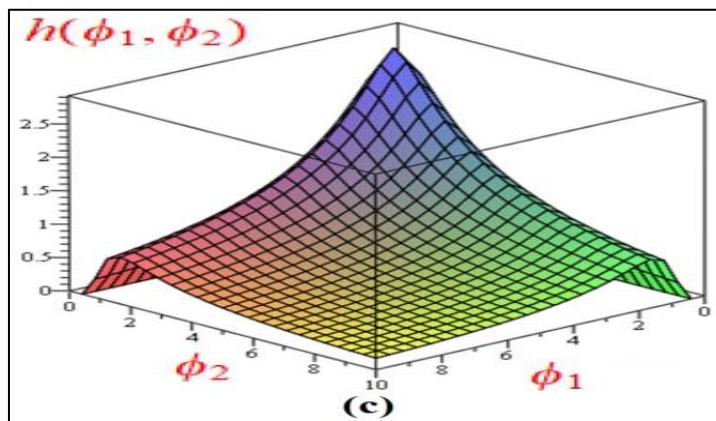
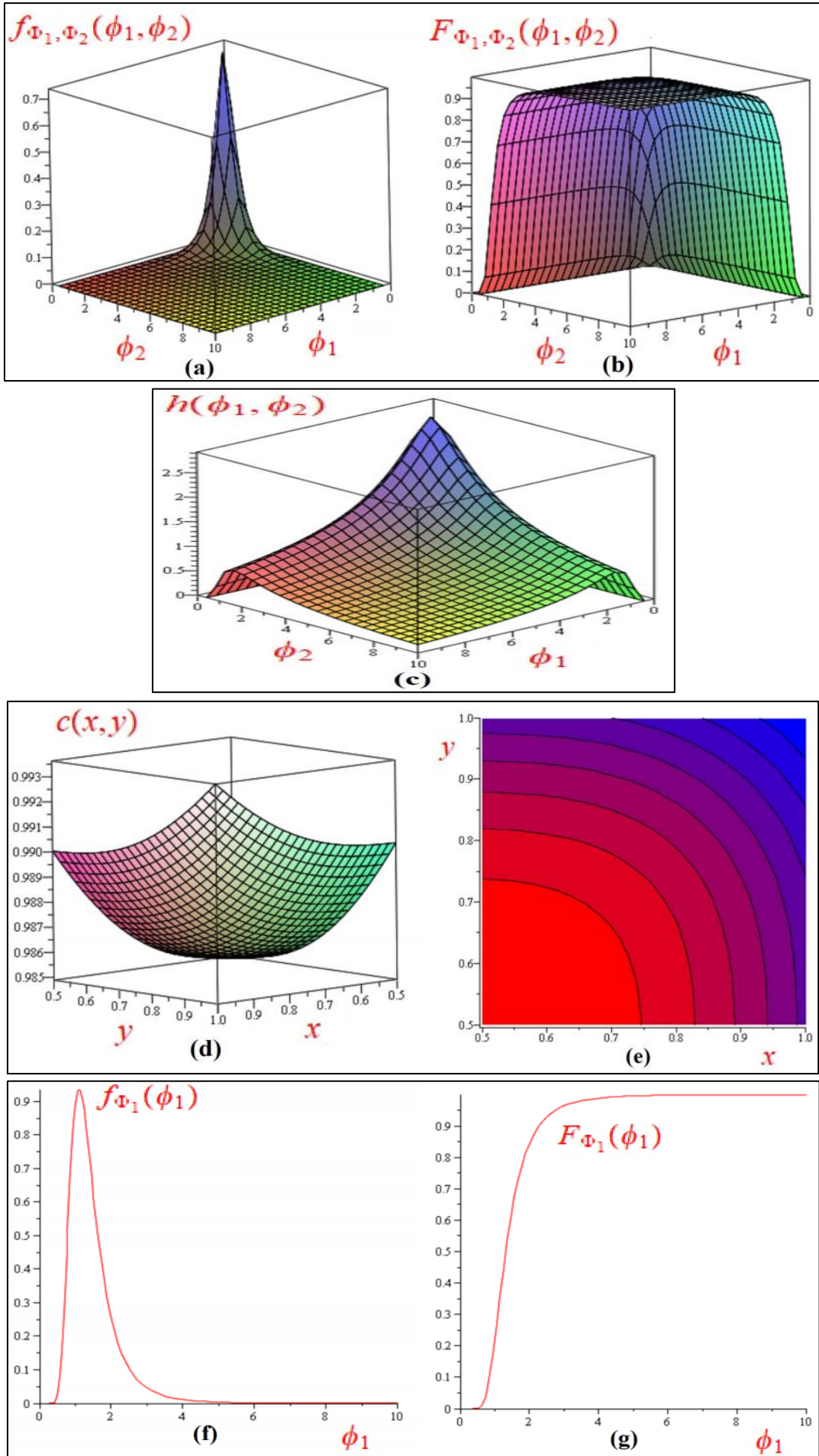


Fig-5. Functions corresponding to $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3) = (0.9, 1.7, 1.9, 1.6, 1.3, 1.6, 0.6, 0.7, 0.8)$.



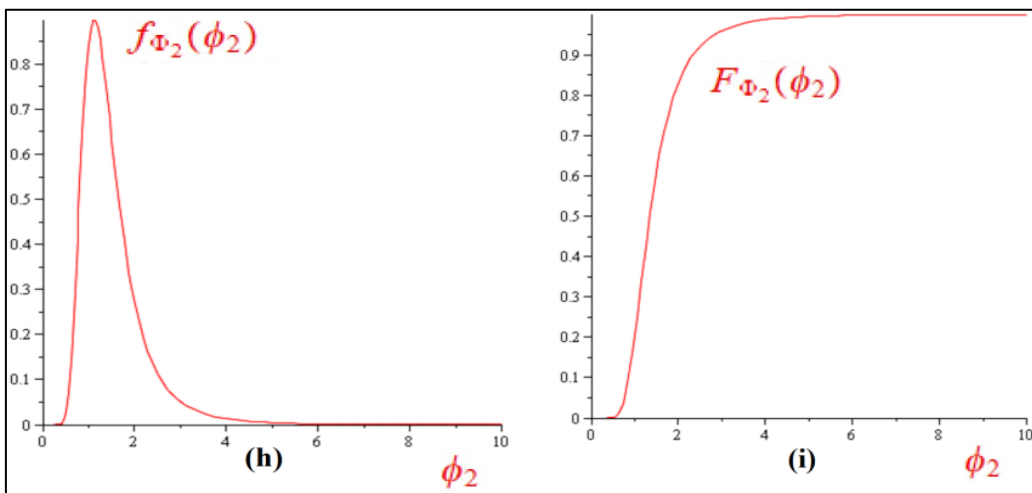


Table 5. Estimated parameters corresponding to $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (3.0, 2.0, 1.0, 4.0, 3.0, 2.0)$.

Sample Size	Parameter	Average Estimate	Bias	SE
n = 30	ϵ	3.53	0.530	1.01
	ε	3.76	1.76	1.02
	ζ	0.378	-0.622	0.0687
	η	4.004	0.004	0.421
	θ	2.958	0.042	0.220
	ϑ	3.21	1.21	1.25
n = 50	ϵ	2.541	-0.495	0.441
	ε	3.37	1.37	1.26
	ζ	0.695	-0.305	0.150
	η	4.344	0.344	0.760
	θ	2.623	-0.377	0.292
	ϑ	2.95	0.950	1.93
n = 120	ϵ	2.854	-0.146	0.387
	ε	2.371	0.371	0.870
	ζ	0.953	-0.047	0.243
	η	4.001	0.001	0.604
	θ	3.342	-0.342	0.372
	ϑ	2.398	0.398	0.453

Table-6. Estimated parameters corresponding to $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (1.4, 1.5, 1.8, 1.4, 1.8, 1.2)$.

Sample Size	Parameter	Average Estimate	Bias	SE
n = 30	ϵ	1.565	0.165	0.273
	ε	2.214	0.714	0.238
	ζ	2.554	0.754	0.659
	η	1.578	0.178	0.244
	θ	1.598	-0.202	0.179
	ϑ	1.484	0.284	0.285
n = 50	ϵ	1.271	-0.129	0.216
	ε	2.009	0.509	0.422
	ζ	2.248	0.448	0.328
	η	1.860	0.460	0.208
	θ	2.156	0.356	0.138
	ϑ	2.977	1.777	0.274
n = 120	ϵ	1.421	0.021	0.249
	ε	1.718	0.218	0.196
	ζ	2.077	0.277	0.296
	η	1.645	0.254	0.169
	θ	1.639	-0.161	0.047
	ϑ	1.284	-0.084	0.225

Table-7. Estimated parameters corresponding to $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta) = (2;5; 1;7; 0;9;2;6; 2;7; 0;8)$.

Sample Size	Parameter	Average Estimate	Bias	SE
n = 30	ϵ	2.002	-0.498	0.501
	ε	2.105	0.405	0.424
	ζ	1.558	0.658	0.423
	η	3.897	1.297	0.901
	θ	2.122	-0.578	0.357
	ϑ	1.378	0.587	0.460
n = 50	ϵ	2.435	-0.065	0.500
	ε	1.748	0.048	0.370
	ζ	1.942	1.042	0.400
	η	3.480	0.880	0.840
	θ	2.504	-0.196	0.302
	ϑ	1.290	0.490	1.470
n = 120	ϵ	2.550	0.050	0.447
	ε	1.793	0.093	0.342
	ζ	0.646	-0.254	0.422
	η	2.973	0.373	0.320
	θ	2.786	0.186	0.245
	ϑ	0.972	0.172	0.474

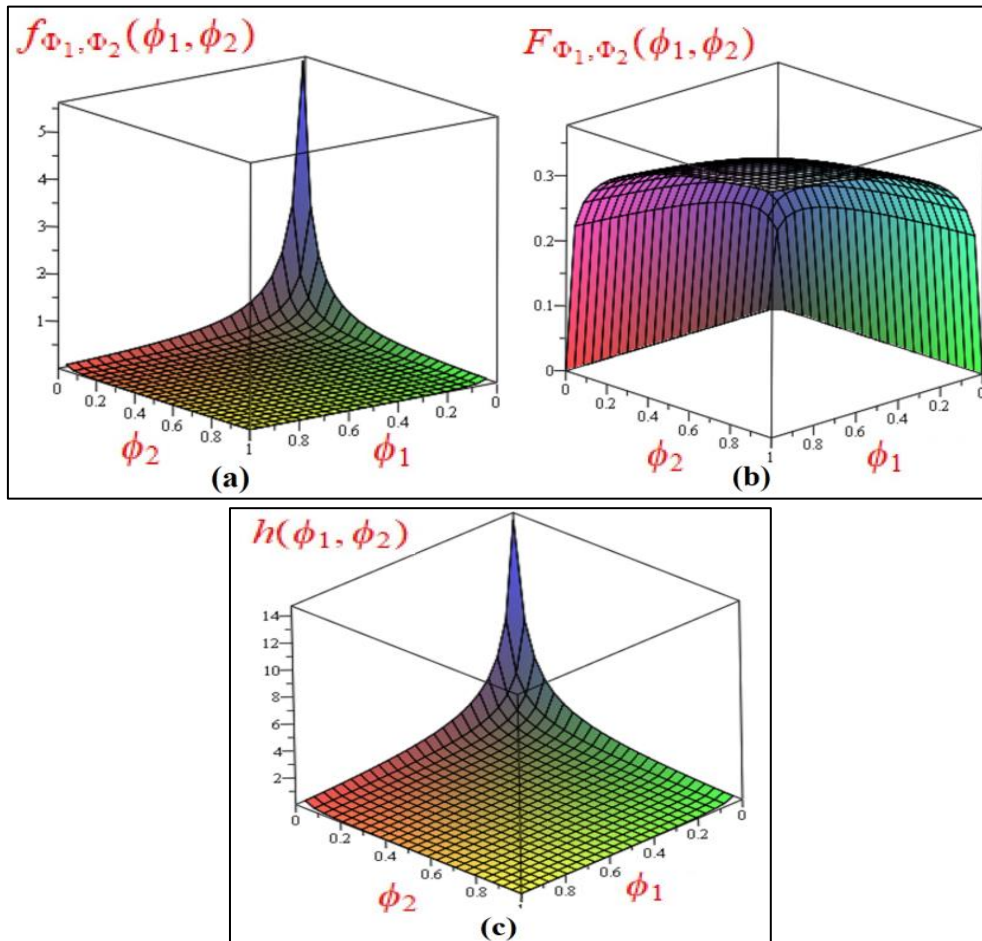
9. Mortality COVID-19 Data

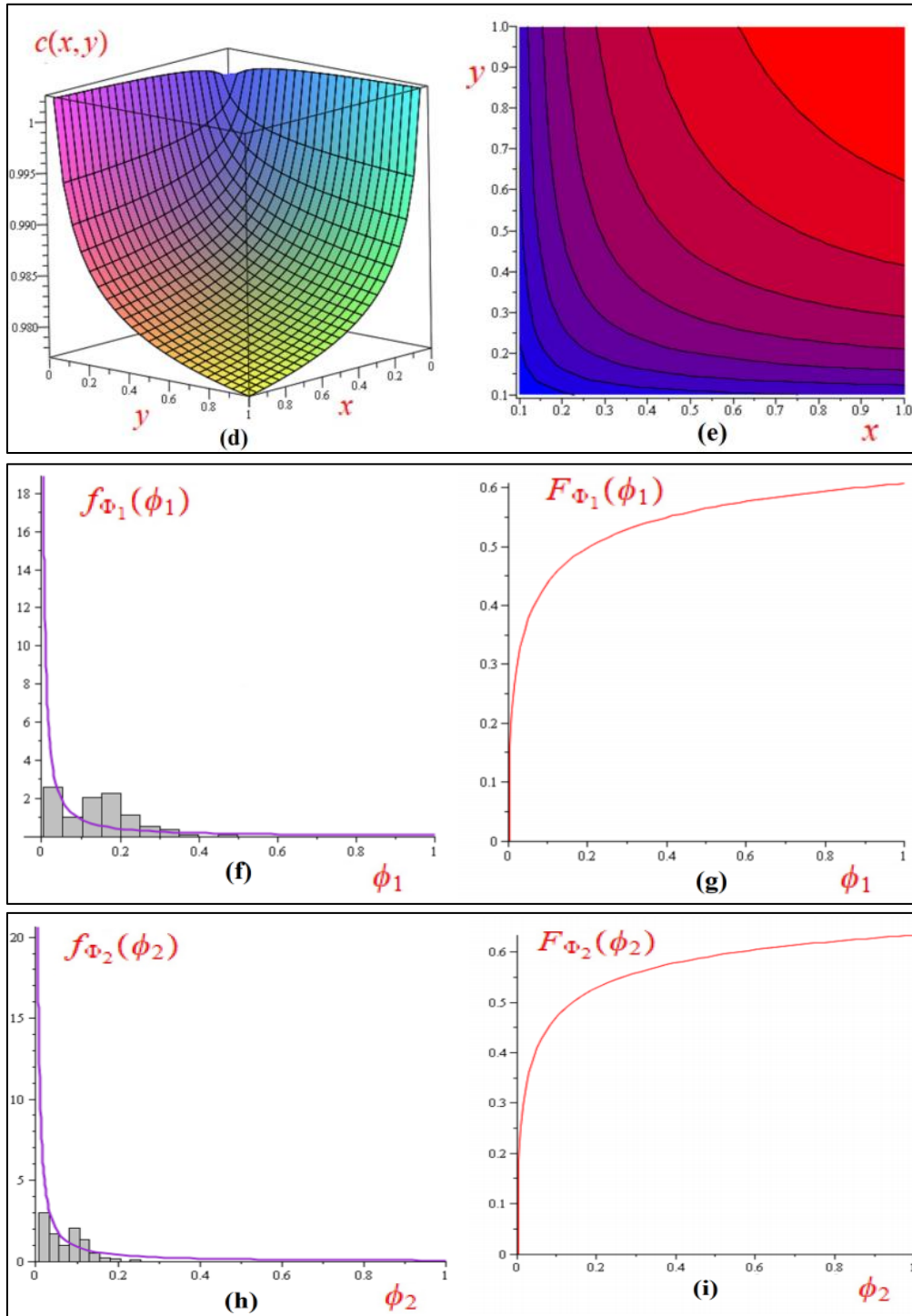
Here the new *BISPTIHLIW* $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$ is used to model the mortality COVID-19 data for Italy and Canada in the period from 1 April to 21 August 2020. Data is available at [https://github.com/CSSEGISandData/ COVID-19/](https://github.com/CSSEGISandData/COVID-19/) and in [32, 33]. We consider the two dimensional random variable (Φ_1, Φ_2) , with observed values the mortality COVID-19 data. The estimated parameters are given in Table 8. Its corresponding functions are plotted in Fig. 6.

Table-8. Estimated parameters for the $(\epsilon, \varepsilon, \zeta, \eta, \theta, \vartheta, \gamma_1, \gamma_2, \gamma_3)$.

Parameter	Estimate	AIC	BIC
ϵ	4.66303844		
ε	0.38218062		
ζ	0.50162001		
η	0.33702847	-133.2667	-106.7279
θ	0.26290366		
ϑ	0.10324411		
γ_1	0.07491866		
γ_2	0.20259496		
γ_3	-0.57018218		

Fig-6. Statistical quantities for the random variable $(U, V) \sim \text{BIETIHLIW}(3.4884764, 0.5391180, 1.0062037, 2.9732696, 10.3121449, 0.2651493, 0.9940035, 0.6297572)$





10. Conclusion

In this study a new univariate six-parameter type I half-logistic inverse Weibull distribution has been introduced, that is a generalization of the model introduced by Elhassanein [18]. Statistical properties of the new model including the pdf, the cdf, the s^{th} moment and the moment generating function have been computed in explicit forms. The flexibility of the model has been proved for different values of parameters. The results of simulation study showed the good performance of the model in terms of maximum likelihood method. The model showed good results via comparison with competitive models and applicability for different types of data. The bivariate extension BISPTIHLIW model has been formulated. That has an absolutely continuous pdf and its statistical quantities are available in explicit forms. Simulation showed a good performance of the model with respect to the goodness of fit. The applicability of the bivariate model has been proved for different types of data using the COVID-19 mortality data for Italy and Canada that are treated as a bivariate random variable.

Conflict of Interest

The author declares that there are no conflicts of interest regarding the publication of this article.

Funding

There is no fund for this article.

References

- [1] Hassan, A. S., 2021. "Kumaraswamy inverted Topp-Leone distribution with applications to COVID-19 data." *CMC*, vol. 86, pp. 337-358.
- [2] Iqbal, Z., Tahir, M. M., and Rizan, N., 2017. "Generalized inverted Kumaraswamy distribution, properties and application." *Open. J. Stat.*, vol. 7, pp. 645-662.
- [3] Ragab, M. and Elhassanein, A., 2022. "A new bivariate extended generalized inverted Kumaraswamy Weibull distribution." *Advances In Mathematical Physics*, vol. 2022, p. 1243018.
- [4] Ramzan, Q., 2022. "On the extended generalized inverted Kumaraswamy distribution." *Computational Intelligence and Neuroscience*, vol. 2022, p. 1612959.
- [5] Ramzan, Q., Amin, M., Elhassanein, A., and Ikram, M., 2021. "The extended generalized inverted Kumaraswamy Weibull distribution: properties and applications." *AIMS Mathematics*, vol. 6, pp. 9955-9980.
- [6] Keller, A. Z. and Kamath, A. R. R., 1982. "Alternative Reliability Models for Mechanical Systems." In *Proceeding of the 3rd International Conference on Reliability and Maintainability*. pp. 411-415.
- [7] Erto, P. and Rapone, M., 1984. "Non-informative and practical Bayesian confidence bounds for reliable life in the Weibull model." *Reliability Engineering*, vol. 7, p. 181-191.
- [8] Akgul, F. G., Senoglu, B., and Arslan, T., 2016. "An alternative distribution to Weibull for modeling the wind speed data: Inverse Weibull distribution." *Energy Conversion and Management* vol. 114 pp. 234-240.
- [9] De Gusmo, F. R. S., Ortega, E. M. M., and Cordeiro, G. M., 2009. "The generalized inverse Weibull distribution." *Stat Papers*, Available: [10.1007/s00362-009-0271-3](https://doi.org/10.1007/s00362-009-0271-3)
- [10] Oluyede, B. O. and Yang, T., 2014. "Generalizations of the inverse Weibull and related distributions with applications." *J. Appl. Stat. Anal*, vol. 7, pp. 94-116.
- [11] Jana, N. and Bera, S., 2022. "Interval estimation of multicomponent stress strength reliability based on inverse Weibull distribution." *Mathematics and Computers in Simulation* 191, pp. 95-119.
- [12] Okasha, H. M., El-Baz, A. H., Tarabiac, A. M. K., and Basheer, A. M., 2017. "Extended inverse Weibull distribution with reliability application." *Journal of the Egyptian Mathematical Society*, vol. 25, pp. 343-349.
- [13] Cordeiro, G. M., Alizadeh, M., Diniz, P. R., and Marinho, P. R. D., 2016. "The type I half-logistic family of distributions." *Journal of Statistical Computation and Simulation*, vol. 86, pp. 707-728.
- [14] Alkarni, S., AOfy, A. Z., Elbatal, I., and Elgarhy, M., 2020. "The extended inverse Weibull distribution: properties and applications." *Complexity*, p. 3297693.
- [15] AL-Moisheer, A. S., 2020. "Bivariate mixture of inverse Weibull distribution: properties and estimation." *Mathematical Problems in Engineering*, vol. 2020, p. 5234601.
- [16] Balakrishnan, N. and Lai, C., 2009. *Continuous bivariate distributions*. 2ed. Springer.
- [17] Darwish, J. A., Al Turk, L. I., and Shahbaz, M. Q., 2021. "Bivariate transmuted Burr distribution: Properties and applications." *Pak. J. Stat. Oper. Res.*, vol. 17, pp. 15-24.
- [18] Elhassanein, A., 2022. "On statistical properties of a new bivariate modified Lindley distribution with an application to financial data." *Complexity*, vol. 2022, p. 2328831.
- [19] Ganji, M., Bevraniand, H., and Golzar, H., 2018. "A new method for generating a continuous bivariate distribution families." *JIRSS*, vol. 17, pp. 109-129.
- [20] Gross, A. J. and Clark, V. A., 1975. *Survival Distributions. Reliability Applications in the Biomedical Sciences*. New York: John Wiley and Sons. p. 1975.
- [21] Kundu, D. and Gupta, A., 2017. "On Bivariate Inverse Weibull Distribution." *Brazilian Journal of Probability and Statistics*, vol. 31, pp. 275-302.
- [22] Kundu, D. and Gupta, R. D., 2009. "Bivariate generalized exponential distribution." *J. Multivariate Anal.*, vol. 100, pp. 581-593.
- [23] Kundu, D. and Gupta, R. D., 2010. "A class of bivariate models with proportional reversed hazard marginals." *Sankhya B* vol. 72, pp. 236-253.
- [24] Mondal, S. and Kundu, D., 2021. "A bivariate inverse Weibull distribution and its applications in complementary risks models." *Journal of Applied Statistics*, vol. 47, pp. 1084-1108.
- [25] Muhammed, H. Z., 2017. "Bivariate dagum distribution." *Int. J. Reliab. Appl.*, vol. 18, pp. 65-82.
- [26] Muhammed, H. Z., 2019. "Bivariate generalized burr and related distributions: Properties and estimation." *J. Data Sci.*, vol. 17, pp. 532-548.
- [27] Thomas, P. Y. and Jose, J., 2020. "A new bivariate distribution with Rayleigh and Lindley distributions as marginals." *Journal of Statistical Theory and Practice* vol. 14, p. 28.
- [28] Vaidyanathan, V. S. and Varghese, A. S., 2016. "Morgenstern type bivariate Lindley distribution. Stat., Optim." *Inf. Comput*, vol. 4, pp. 132-146.
- [29] AOfy, A. Z., 2017. "The odd Exponentiated half-logistic-G family: properties, characterizations and applications." *Chilean Journal of Statistics*, vol. 8, pp. 65-91.
- [30] Durrani, T. S. and Zeng, X., 2007. "Copulas for bivariate probability distributions." *Electronics Letters*, vol. 43, pp. 248-249.
- [31] Navarro, J., 2010. "Characterizations using the bivariate failure rate function." *Statistics and Probability Letters*, vol. 78 p. 1349.

- [32] Algarni, A., 2021. "Type I half logistic Burr X-G family: Properties, bayesian, and non-Bayesian estimation under censored samples and applica-tions to COVID-19 data." *Mathematical Problems in Engineering*, vol. 2021, p. 5461130.
- [33] Bantan, R. A. R., 2022. "Statistical analysis of COVID 19 data: Using a new univariate and bivariate statistical model." *Journal of Function Spaces*, vol. 2022, p. 2851352.